

# Bayesian parametric models to handle missing longitudinal outcome data in trial-based health economic evaluations

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Joint work with **Andrea Gabrio (UCL)** and **Michael J Daniels (University of Florida)**

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Invited session: "Sensitivity analysis for uncheckable assumptions"

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## 1. Health economic evaluation

- What is it?
- How does it work?

## 2. Statistical modelling

- Standard approach
- The importance of being a Bayesian

## 3. Bayesian modelling for missing data in HTA

- Modelling & advantages
- Bayesian nature of dealing with missing data

## 4. Motivating example

- Data (and their weird features...) & Bayesian modelling
- Results

## 5. Conclusions

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## Statistical model

- Estimates relevant **population** parameters  $\theta$
- Varies with the type of available data (& statistical approach!)



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$$\Delta_e = f_e(\theta)$$

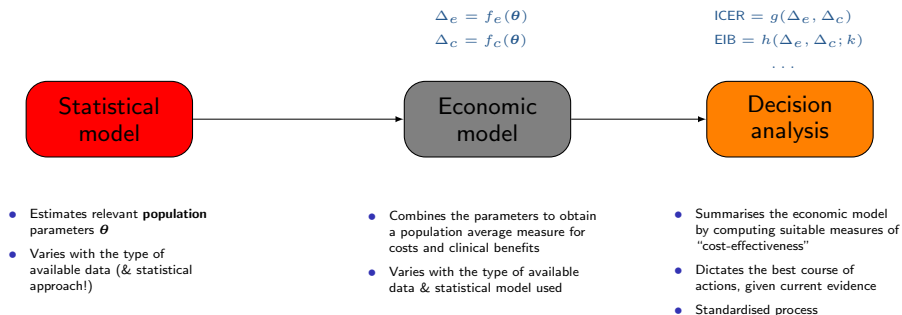
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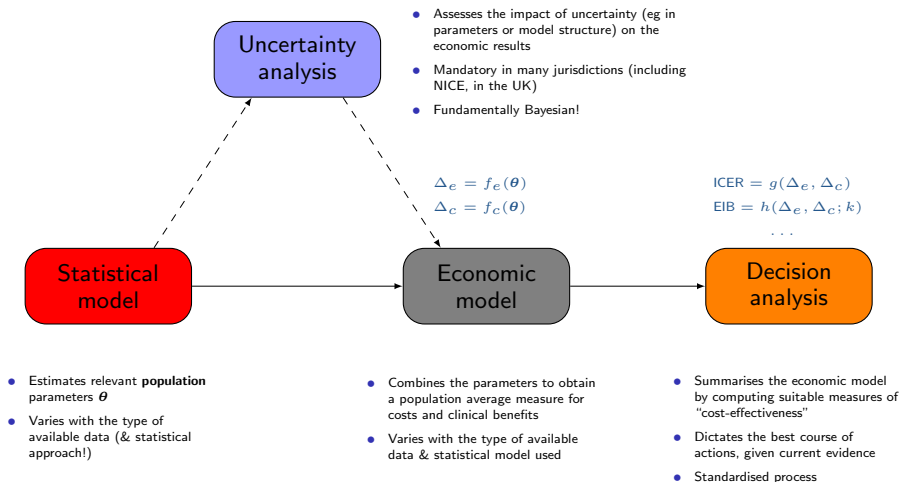
- Estimates relevant **population** parameters  $\theta$
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- Combines the parameters to obtain a population average measure for costs and clinical benefits
- Varies with the type of available data & statistical model used

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ID	Trt	Demographics			HRQL data				Resource use data				Clinical outcome			
		Sex	Age	...	$u_0$	$u_1$	...	$u_J$	$c_0$	$c_1$	...	$c_J$	$y_0$	$y_1$	...	$y_J$
1	1	M	23	...	0.32	0.66	...	0.44	103	241	...	80	$y_{10}$	$y_{11}$	...	$y_{1J}$
2	1	M	21	...	0.12	0.16	...	0.38	1204	1808	...	877	$y_{20}$	$y_{21}$	...	$y_{2J}$
3	2	F	19	...	0.49	0.55	...	0.88	16	12	...	22	$y_{30}$	$y_{31}$	...	$y_{3J}$
...	...	...	...	...	...	...	...	...	...	...	...	...	...	...	...	...

$y_{ij}$  = Survival time, event indicator (eg CVD), number of events, continuous measurement (eg blood pressure), ...

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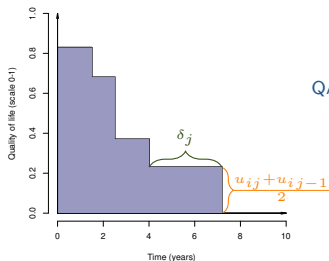
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1 Compute individual QALYs and total costs as

$$e_i = \sum_{j=1}^J (u_{ij} + u_{i,j-1}) \frac{\delta_j}{2} \quad \text{and} \quad c_i = \sum_{j=0}^J c_{ij}, \quad \left[ \text{with: } \delta_j = \frac{\text{Time}_j - \text{Time}_{j-1}}{\text{Unit of time}} \right]$$



$QALY_i = \text{“Area under the curve”}$

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2 (Often implicitly) assume normality and linearity and model **independently** individual QALYs and total costs by controlling for baseline values

$$\begin{aligned} e_i &= \alpha_{e0} + \alpha_{e1}u_{0i} + \alpha_{e2}\text{Trt}_i + \varepsilon_{ei} [+ \dots], & \varepsilon_{ei} &\sim \text{Normal}(0, \sigma_e) \\ c_i &= \alpha_{c0} + \alpha_{c1}c_{0i} + \alpha_{c2}\text{Trt}_i + \varepsilon_{ci} [+ \dots], & \varepsilon_{ci} &\sim \text{Normal}(0, \sigma_c) \end{aligned}$$

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3 Estimate population average cost and effectiveness differentials and use bootstrap to quantify uncertainty

## What's wrong with this?...

- Potential correlation between costs & clinical benefits [Individual Level + Aggregated Level Data]
  - Strong positive correlation — effective treatments are innovative and result from intensive and lengthy research  $\Rightarrow$  are associated with higher unit costs
  - Negative correlation — more effective treatments may reduce total care pathway costs e.g. by reducing hospitalisations, side effects, etc.
  - Because of the way in which standard models are set up, bootstrapping generally only approximates the underlying level of correlation — **MCMC does a better job!**

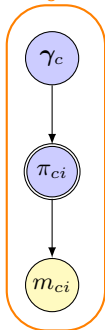
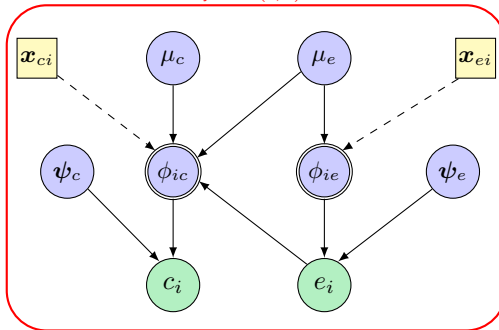
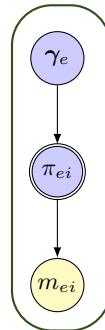


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- Joint/marginal normality not realistic [Mainly ILD]
  - Costs usually skewed and benefits may be bounded in  $[0; 1]$
  - Can use transformation (e.g. logs) — but care is needed when back transforming to the natural scale
  - Should use more suitable models (e.g. Beta, Gamma or log-Normal) — **generally easier under a Bayesian framework**
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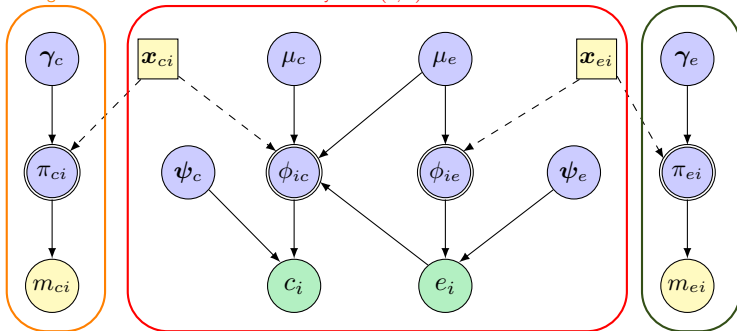
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- ... and of course **Partially Observed** data
  - Can have item and/or unit non-response
  - Missingness may occur in either or both benefits/costs
  - The missingness mechanisms may also be correlated
  - Focus in decision-making, not inference — **Bayesian approach particularly suited for this!**

MCAR ( $e, c$ )Model of missingness for  $c$ Model of analysis for ( $c, e$ )Model of missingness for  $e$ 

- Partially observed data
- Unobservable parameters
- ◎ Deterministic function of random quantities
- Fully observed, unmodelled data
- Fully observed, modelled data

- $m_{ei} \sim \text{Bernoulli}(\pi_{ei}); \quad \text{logit}(\pi_{ei}) = \gamma_{e0}$
- $m_{ci} \sim \text{Bernoulli}(\pi_{ci}); \quad \text{logit}(\pi_{ci}) = \gamma_{c0}$

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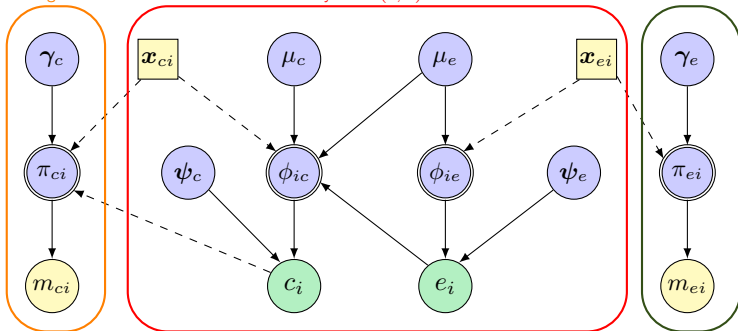
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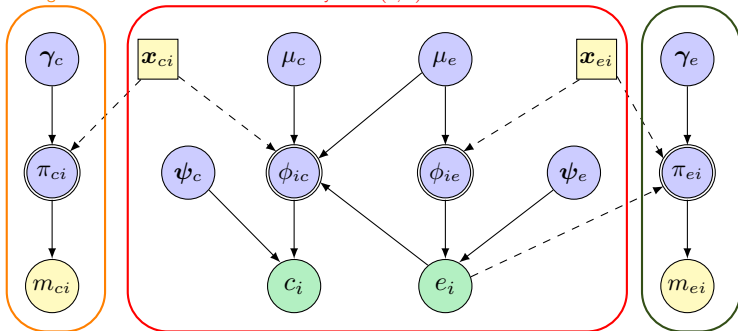
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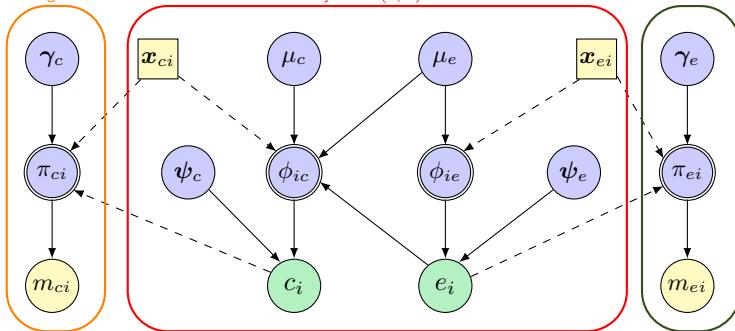
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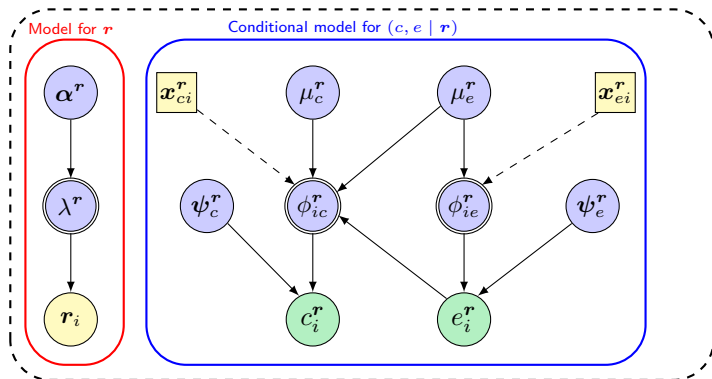
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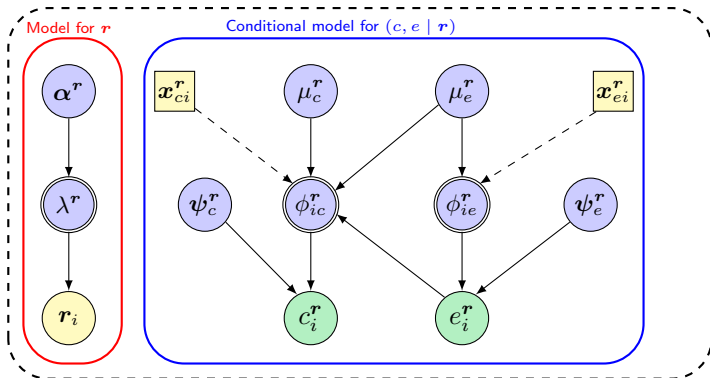
MAR ( $e, c$ )

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For  $\mathbf{r} \in \mathcal{R} \equiv [(1, 1), (0, 1), (1, 0), (0, 0)]$ 

- $\mathbf{r}_i = (m_{ei}, m_{ci}) \sim \text{Multinomial}(\boldsymbol{\lambda}^{\mathbf{r}})$ ;  $\mu_e = \sum_{\mathbf{r} \in \mathcal{R}} \mu_e^{\mathbf{r}} \lambda^{\mathbf{r}};$   $\mu_c = \sum_{\mathbf{r} \in \mathcal{R}} \mu_c^{\mathbf{r}} \lambda^{\mathbf{r}}$
- CC restriction:  $\mu_e^{(0,1)} = \mu_e^{(0,0)} = \mu_e^{(1,1)};$   $\mu_c^{(1,0)} = \mu_c^{(0,0)} = \mu_c^{(1,1)}$

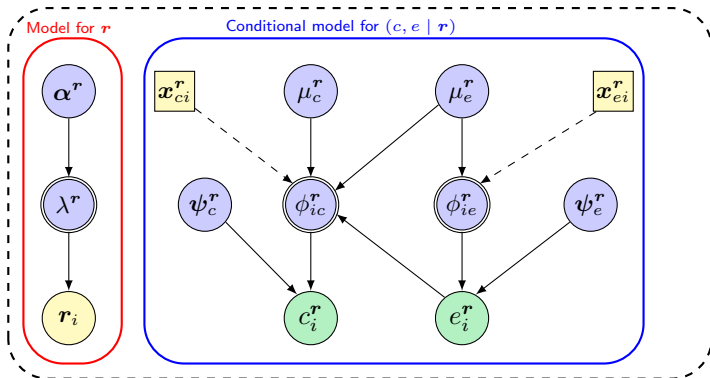


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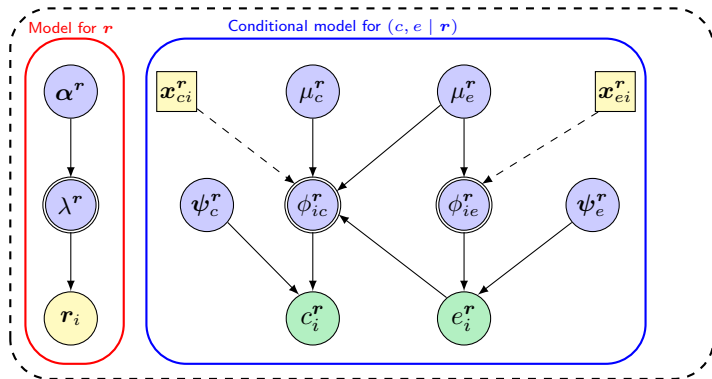
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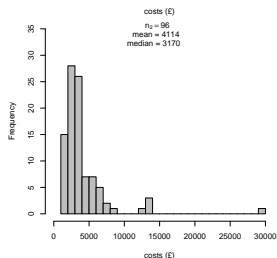
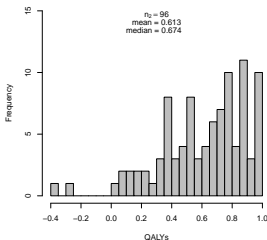
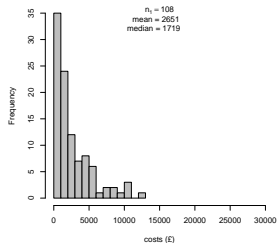
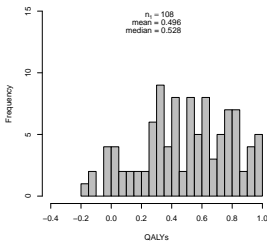
## Example: PBS trial

- Multi-centre RCT that evaluates the cost-effectiveness of a new multicomponent intervention for individuals suffering from intellectual disability and challenging behaviour
- Both utilities (EQ-5D) and costs (clinic records) are partially-observed

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Time	Control ( $n_1=136$ )		Intervention ( $n_2=108$ )	
	observed (%)		observed (%)	
	utilities	costs	utilities	costs
Baseline	127 (93%)	136 (100%)	103 (95%)	108 (100%)
6 months	119 (86%)	128 (94%)	102 (94%)	103 (95%)
12 months	125 (92%)	130 (96%)	103 (95%)	104 (96%)
<b>complete cases</b>	108 (79%)		96 (89%)	

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- In reality, the data have a longitudinal nature and particularly in the presence of missing data we have several advantages in fully exploiting it
  - Account for time dependence between outcomes  $\mathbf{y}_{ij} = (u_{ij}, c_{ij})$
  - Use all available utility/cost data in each pattern  $\mathbf{r}_{ij} = (r_{ij}^u, r_{ij}^c)$

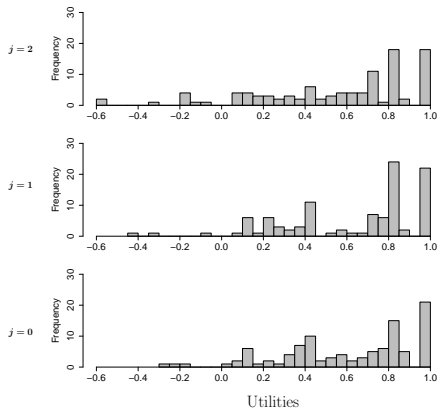
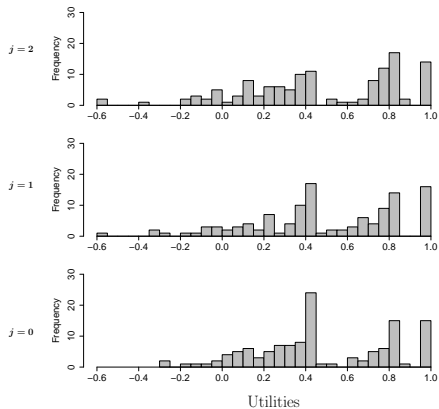
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  - Account for time dependence between outcomes  $\mathbf{y}_{ij} = (u_{ij}, c_{ij})$
  - Use all available utility/cost data in each pattern  $\mathbf{r}_{ij} = (r_{ij}^u, r_{ij}^c)$
- Can use **pattern mixture model**
  - 1 Factorise  $p(\mathbf{y}, \mathbf{r})$  into  $p(\mathbf{y}_{obs}^r, \mathbf{r})$  and  $p(\mathbf{y}_{mis}^r | \mathbf{y}_{obs}^r, \mathbf{r})$
  - 2 Integrate out  $\mathbf{y}_{mis}^r$  from  $p(\mathbf{y}, \mathbf{r})$  and estimate the means of  $\mathbf{y}_{obs}^r$
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- Assess the robustness of the results to plausible MNAR scenarios using different informative priors on  $\Delta$

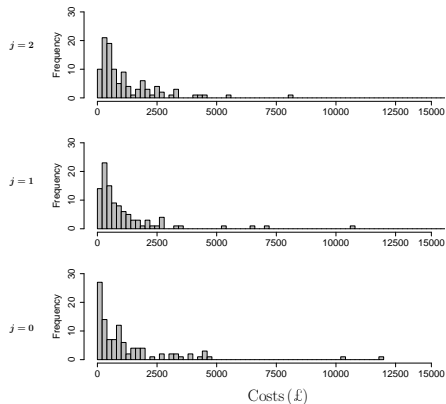
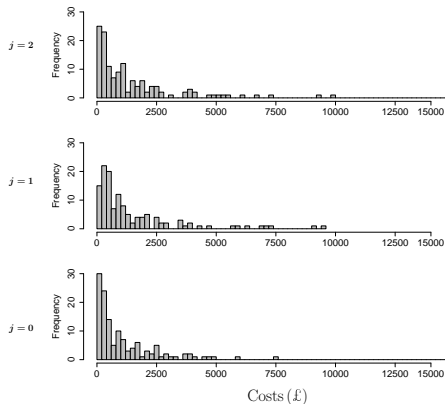
## Control

## Intervention



## Control

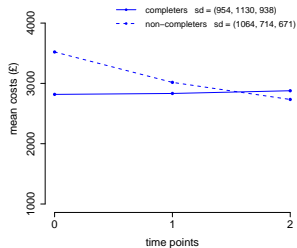
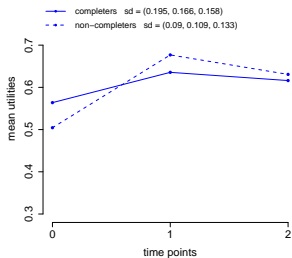
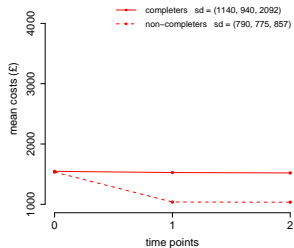
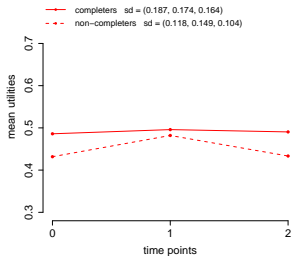
## Intervention



		Control ( $t = 1$ )						$n_1^r$	Intervention ( $t = 2$ )						$n_2^r$
		$u_0$	$c_0$	$u_1$	$c_1$	$u_2$	$c_2$		$u_0$	$c_0$	$u_1$	$c_1$	$u_2$	$c_2$	
$r$		1	1	1	1	1	1	<b>108</b>	1	1	1	1	1	1	<b>96</b>
mean		0.678	1546	0.684	1527	0.680	1520		0.726	2818	0.771	2833	0.759	2878	
$r$		0	1	1	1	1	1	<b>7</b>	0	1	1	1	1	1	<b>5</b>
mean		-	1310	0.704	1440	0.644	1858		-	2573	0.780	2939	0.849	2113	
$r$		1	1	0	1	1	1	<b>4</b>	1	1	0	1	1	1	<b>1</b>
mean		0.709	1620	-	1087	0.737	851		0.467	9649	-	4828	0.259	4930	
$r$		1	1	1	1	0	1	<b>2</b>	1	1	1	1	0	1	<b>1</b>
mean		0.564	640	0.648	512	-	286		0.817	3788	0.884	0	-	0	
$r$		1	1	0	0	1	1	<b>4</b>	1	1	0	0	1	1	<b>1</b>
mean		0.716	2834	-	-	0.634	679		0.501	3608	-	-	0.872	4781	
$r$		1	1	0	0	0	0	<b>4</b>	1	1	0	0	0	0	<b>4</b>
mean		0.434	1528	-	-	-	-		0.760	3086	-	-	-	-	
$r$		0	1	0	1	1	1	<b>2</b>	0	1	0	1	1	1	<b>0</b>
mean		-	595	-	397	0.483	69		-	-	-	-	-	-	
$r$		1	1	1	1	0	0	<b>2</b>	1	1	1	1	0	0	<b>0</b>
mean		0.743	1434	0.705	1606	-	-		-	-	-	-	-	-	
$r$		1	1	0	1	0	1	<b>3</b>	1	1	0	1	0	1	<b>0</b>
mean		0.726	1510	-	432	-	976		-	-	-	-	-	-	

$\rightarrow r = 1$

$r \neq 1$



- Fit model to completers  $r = 1$  and joint set of all other patterns  $r \neq 1$  separately for  $t = 1, 2$

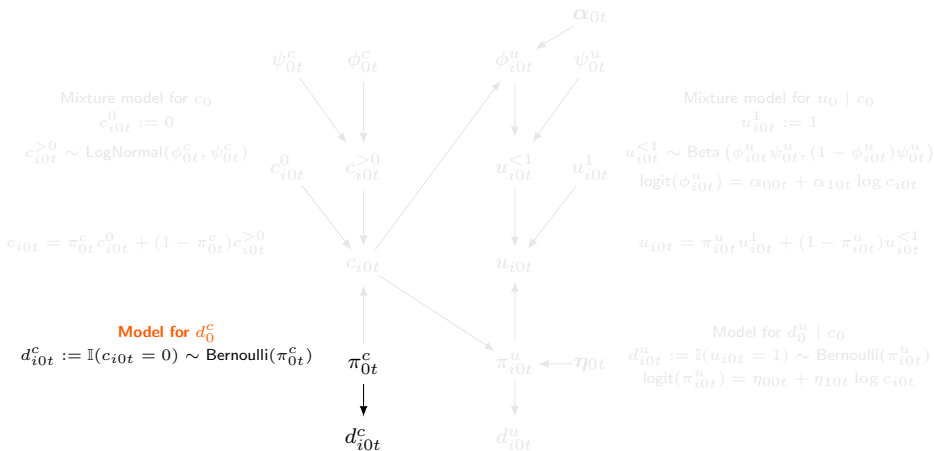
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- Account for skewness using
  - **Beta** distributions for  $u_{ij}^* = \frac{u_{ij} - \min(\mathbf{u}_j)}{\max(\mathbf{u}_j) - \min(\mathbf{u}_j)}$
  - **LogNormal** distributions for  $c_{ij}$

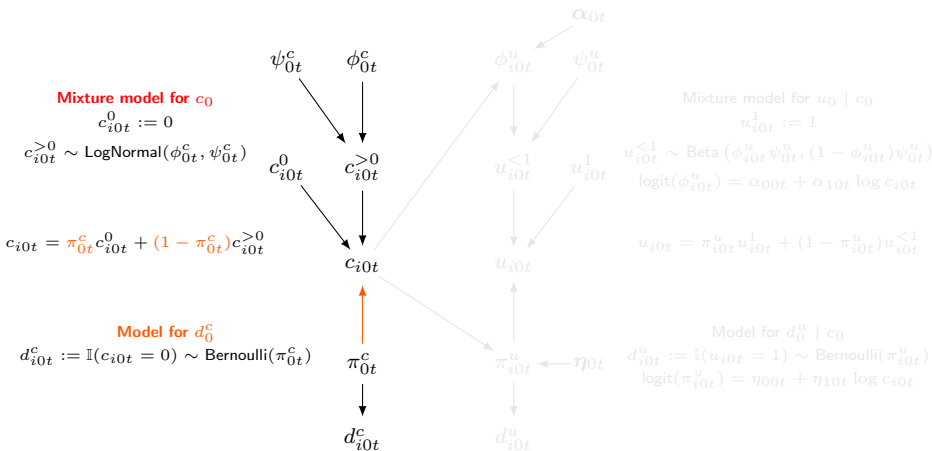


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  - **LogNormal** distributions for  $c_{ij}$
- Allow for structural **ones** in  $u_{ij}$  and **zeros** in  $c_{ij}$ 
  - Define  $d_{ij}^u := \mathbb{I}(u_{ij} = 1)$  and  $d_{ij}^c := \mathbb{I}(c_{ij} = 0)$
  - Use a hurdle model to account for mixture of patients within the groups

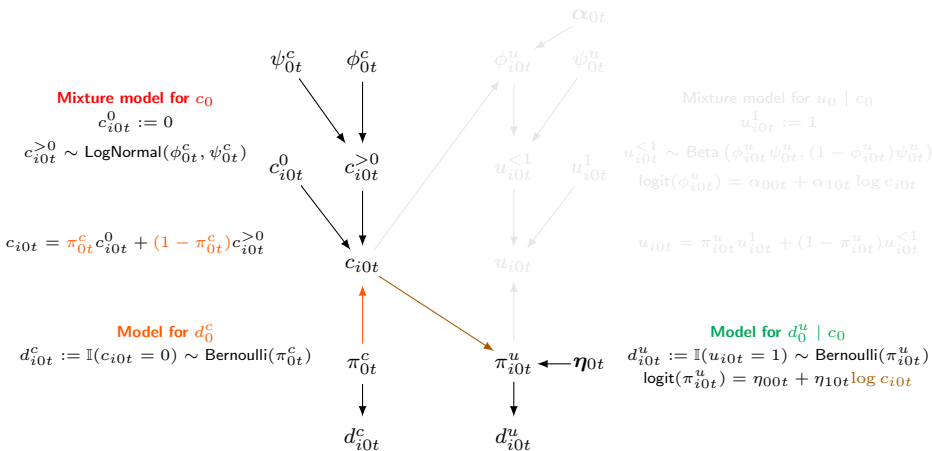
- At  $j = 0$



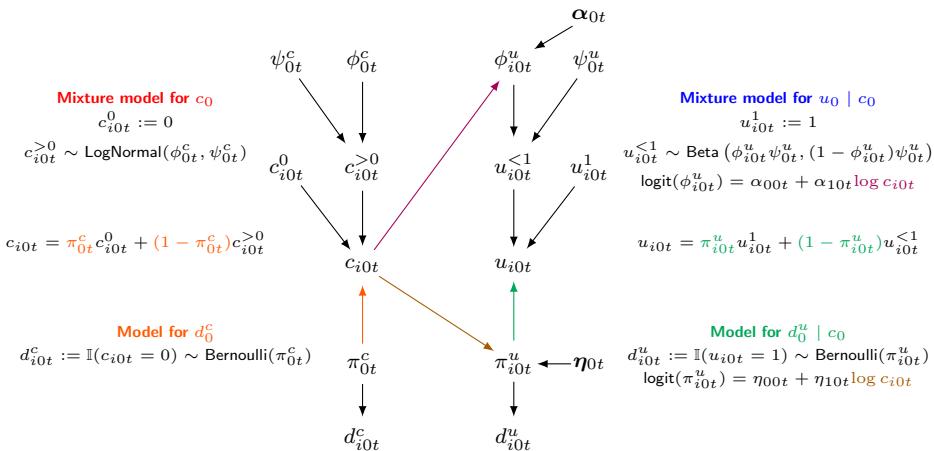
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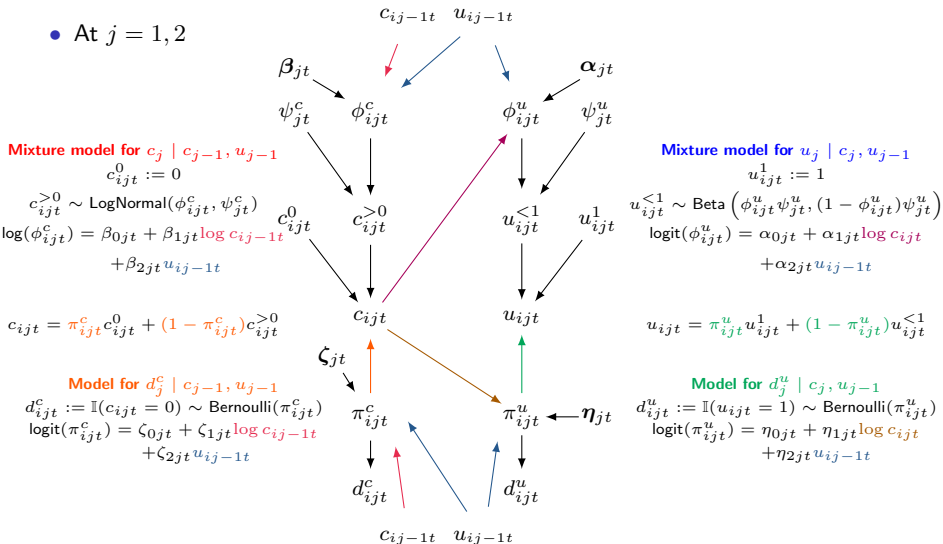
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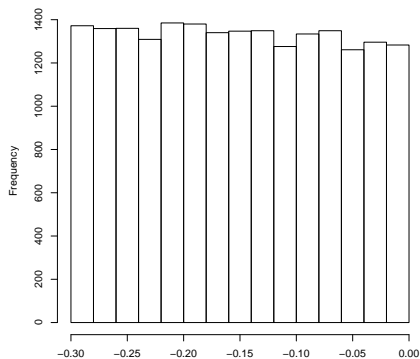


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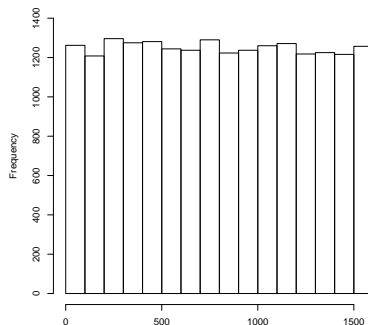
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- Set  $\Delta_j = \mathbf{0}$  as benchmark assumption
- Specify three alternative priors on  $\Delta_j = (\Delta_j^u, \Delta_j^c)$ , calibrated based on the variability in the observed data at each time  $j$

- Assumption:  $u_{mis} < u_{obs}$  and  $c_{mis} > c_{obs}$
- $\Delta^{flat}$ : Flat between 0 and twice the observed standard deviation



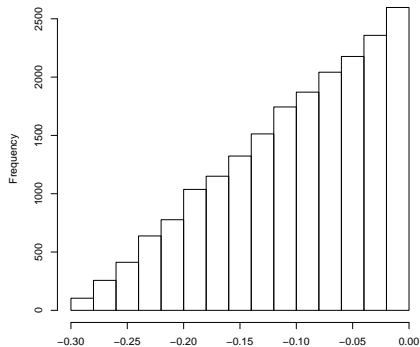
$\Delta_1^u$



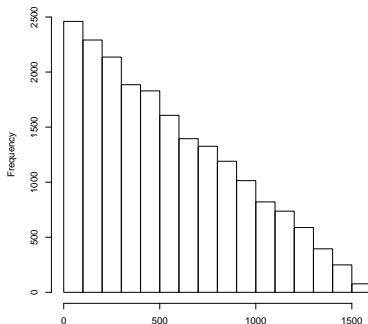
$\Delta_1^c$

# Priors on sensitivity parameters

- Assumption:  $u_{mis} < u_{obs}$  and  $c_{mis} > c_{obs}$
- $\Delta^{skew0}$ : Skewed towards values closer to 0 on the same range as  $\Delta^{flat}$



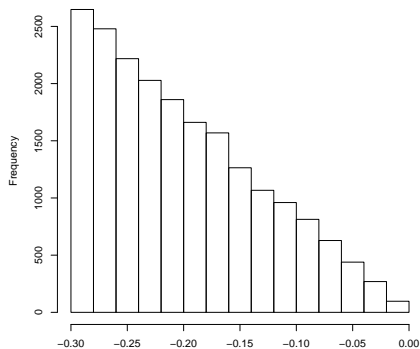
$\Delta_1^u$



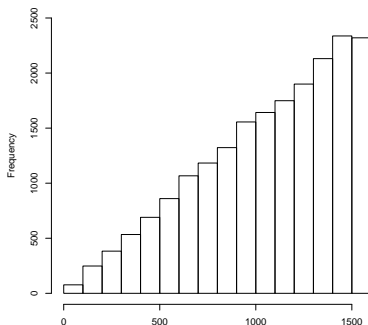
$\Delta_1^c$

# Priors on sensitivity parameters

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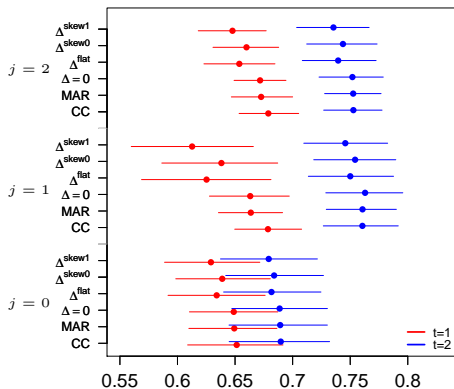
$\Delta_1^u$



$\Delta_1^c$

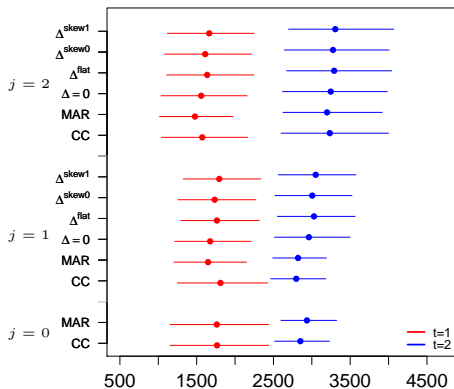
# Results: means utilities and costs

$$\mu_{jt}^u$$



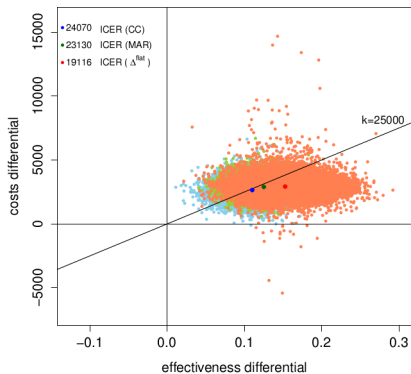
Utilities

$$\mu_{jt}^c$$

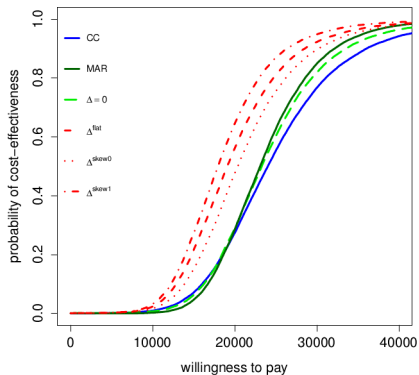


Costs (£)

## Cost-Effectiveness Plane



## Cost-Effectiveness Acceptability Curve





### ④ Flexibility of the modelling framework

- Naturally allows the propagation of uncertainty to the economic model
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- 3 Principled incorporation of **external evidence** through priors
  - Crucial for conducting sensitivity analysis to MNAR
  - Useful in small/pilot trials where there is limited evidence

# Thank you!